

# Ideological Change and the Economics of Voting Behavior in the US, 1920-2008 \*

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## **Abstract**

This paper tests the proposition that voters advance a more liberal agenda in prosperous times and shift towards being more conservative in dire economic times. A reference-dependent utility model establishes a dynamic relationship between income growth and voting behavior by way of the demand for public goods and the optimal tax rate. This income growth model may serve as an alternative to the redistributive model of voter preferences. Ordinary and instrumented statistical analyses of a new time series for the US median voter are encouraging of the theory. This work links voting behavior to economic business cycles and shows that ideological change is endogenous to income growth rates.

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# 1 Introduction

Does a prospering economy push aggregate voter preferences towards more expansionary government and the liberal left; and does a contracting economy lead voters to favor smaller government and the conservative right? In short, is ideological change endogenous to income growth rates?

Surprisingly few scholars have dealt with this important question head-on. Two reasons may possibly explain this odd lack of attention for what could well be a fundamental dynamic in political economy. First, the pioneering work by Campbell, Converse, Miller and Stokes (1960) and Lipset and Rokkan (1967) established models of voter preferences that are determined by partisan affiliation or class and religious cleavages in society. Later work augmented sociological sources with the ways in which electoral systems (Alesina and Glaeser 2004, Iversen and Soskice 2006) and economic institutions (De Neve 2009a, Iversen and Soskice 2009) shape interests, ideology, and voting behavior. The result of these literatures, however, is a rather static vision of ideological change. For voter ideology to change we would need variation in sociological structures or electoral and economic institutions. Because these variables are either hard to quantify or slow-moving, there have been few attempts at better understanding ideological change.

Equally important may be that virtually all attention for the link between economics and voting behavior has been monopolized by short-term analyses of how economic performance affects incumbency. Hence, from the outset, it is important to distinguish the empirical and modeling effort in this paper from the large literature denoted as “economic voting.” What is known as the economic voting literature considers sociotropic and egocentric (or pocketbook) economic effects on incumbent government approval and election outcomes (Kramer 1971, Lewis-Beck 1988, Alesina, Roubini and Cohen 1997, Lewis-Beck and Paldam 2000, Duch and Stevenson 2008). This paper, however, studies how changing economic realities alter political ideology and voting behavior over multiple elections regardless of incumbency. In doing so, it adds to a small number of prominent pieces that have also attempted to capture

this fundamental dynamic in political economy. No previous attempt, however, has been able to cover as much ground, nor had access to the measure of voting behavior presented here. Because swings in ideology happen slowly, the importance of having multi-generational time series cannot be overstated. This emphasis on long time series and ideology distinguishes this research effort from the typical study in economic voting that considers how economic conditions affect an incumbents chances for reelection. While the term economic voting is generic enough to also apply to this paper, it would have to be considered in a category distinct from the aforementioned accountability literature that has come to embody the notion of economic voting.

It is also worthwhile noting from the outset that this is not a theory about how rich or poor Americans vote. This work is preoccupied with how changing economic realities affect aggregate voting behavior; this paper thus hopes to gain understanding into the drivers of ideological change across the US electorate. In his seminal piece “What moves policy sentiment?” Robert Durr (1993) was the first to squarely tackle this deviously simple question. According to Durr, shifts in US domestic policy sentiment on the liberal-conservative spectrum were a response to changing economic expectations. With expectations of a strong economy producing greater support for liberal policies and declining economic conditions shifting the policy mood to the right. Durr’s empirical analysis of US policy sentiment revealed the existence of such ideological undercurrents regardless of incumbency effects. In another prominent piece, Stevenson (2001) expanded on Durr’s theory and conducted a comparative study of fourteen Western democracies to also find that changes in aggregate voter preferences relate systematically to national economic performance. In similar vein, Kim and Fording (2001), Markussen (2008), and Kayser (2007, 2009) explore the interaction between economic conditions and electoral choice in a comparative setting and also point to international economic sources of these seemingly domestic processes. Durr (1993), however, was both the first and the last to take a close empirical look at whether the ideology of the US public moves in sync with the domestic business cycle. His analysis covered the years

1968-88, a relatively short period of time to record what Durr himself described as a long-memoried, dynamic equilibrium between the economy and policy mood that moves in long waves through time. The dependent variable in Durr’s analysis was the notion of “policy sentiment,” a measure devised by Stimson (1991) that aggregates hundreds of distinct US public opinion surveys dealing with a multitude of different policy preferences. When pooled together, Durr and Stimson argued, it becomes possible to construct a single time series for the policy mood that gauges movement along a liberal-conservative continuum.

This paper expands and tests the theory originally proposed by Durr and hopes to place the economics of voting behavior at the heart of political economy. In the process, this work introduces a new time series to capture actual voting behavior, formalizes an income growth model, and runs empirical tests that attempt to deal with the reverse causality of economic voting. The paper starts with a quick overview of ideological change and voting behavior in the US since 1920 and introduces the median voter as our dependent variable. Next, a utility-based model is discussed and the relationship between income growth and voting behavior is established. An empirical analysis that considers reverse causality is carried out for disposable income growth as well as GDP growth, unemployment, and inflation rates. The results and implications of this research are discussed before the paper concludes.

## **2 Ideological change and voting behavior in the US, 1920-2008**

### **2.1 The median voter data set**

In order to quantify ideological change and voting behavior this paper turns to the concept of the median voter. Ever since the seminal works by Black (1948) and Downs (1957) established the concept of the median voter, or the ideological center of the electorate, it has figured widely across literatures. This paper introduces a new median voter data set that

is detailed in a separate research note (De Neve 2009b). The data employs the statistics provided by the Comparative Manifesto Project<sup>1</sup> (Budge, Klingemann, Volkens, Bara and Tanenbaum 2001, Klingemann, Volkens, Bara, Budge and McDonald 2006) but corrects for stochastic error as done by Benoit, Laver and Mikheylov (2009).<sup>2</sup>

In essence, the position of the median voter is computed from vote shares for the ideologically ranked parties; a methodology spearheaded by Kim and Fording (1998). This is done by first ranking the parties by ideological score for every election through textual analysis of party documents. Then for each party the interval where its supporters are located is tabulated by locating the midpoints between the ideologically neighboring parties. Assuming that voters choose the candidate or party that is ideologically closest to them, a party will attract the votes of those that are part of the interval that surrounds that party. The assumption that voting behavior is an expression of ideological beliefs is common (Coate and Conlin 2004, Mullainathan and Washington 2009). Still, it is important to underscore that this assumption implies a disregard of the part of the electorate that may vote strategically rather than ideologically.<sup>3</sup> Finally, the electoral results for each party at every election are matched to produce the percentage of the electorate that is grouped into each ideological

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<sup>1</sup>The Comparative Manifesto Project (CMP) codified all sentences of every election manifesto to place parties on a left-right scale. The scaling consists in subtracting the sum of percentage references to categories grouped as left from the sum of percentage references to categories grouped as right. The manifesto data is collected such that each statement is assigned to either a pro-left or a pro-right category. Consequently, negative scores represent a generally left position, whereas positive scores are reflective of a right position. Results range between -100 (extreme left) and +100 (extreme right).

<sup>2</sup>Benoit et al. (2009) detail the inherently stochastic processes of manifesto authorship and manifesto coding for which the CMP does not provide error estimates. The absence of estimates of measurement uncertainty in the CMP data is troublesome and lowers the scientific quality of its statistics as well as the research that builds upon it. Treating words as data with error, Benoit et al. (2009) proceed by bootstrapping the analysis of every coded manifesto. By way of these simulations they reconstruct the stochastic processes that generated these political texts. In doing so they are able to estimate degrees of non-systematic error for the thousands of manifestos coded by the CMP. The use of these error estimates allows for better empirical and theoretical inferences from the CMP data. The bootstrapping work by Benoit et al (2009) also allows for generating new data estimates of party policy positions. The alternative estimate for a party policy position then becomes the mean estimator of the 1,000 bootstrap simulations that were drawn for each manifesto. This new data for party policy positions calibrates for stochastic error in the CMP. De Neve (2009b) adopts these corrected party policy positions in order to construct a new median voter data set.

<sup>3</sup>Stevenson (2001) notes that estimates of the importance of strategic voting rarely attain 10%.

interval. As Kim and Fording (2003) point out, this methodology requires us “to conceive of elections as large-scale opinion polls.” Where the ballot acts as a survey in which the subject chooses the party that is ideologically closest on the partisan left-right spectrum. As such, it is possible to treat election results as a grouped frequency distribution and tabulate a median statistic. The results range between -100 (extreme left) and +100 (extreme right) and non-election values are interpolated linearly.<sup>4</sup> De Neve (2009b) provides details on the precise tabulations and assumptions that were involved to build these new median voter statistics and extends the data to include 25 Western democracies. It is important to highlight that the median voter statistics are derived indirectly via party policy positions and their success at the election polls, rather than direct evidence of voter opinions. Notwithstanding these limitations, the median voter data presented here are unique in their range and detail. Moreover, as compared to self-placement surveys, this methodology ties in actual voting behavior which leaves opinion polling data more appropriately coined as “median citizen” instead of “median voter”.

The methodology to construct these median voter statistics underwent robustness checks carried out by, among others, Powell (2000) and McDonald and Budge (2005). Table 1 compares the US median voter data used in this paper with alternative measures from the American National Elections Studies (ANES), Kim and Fording (2003), and Ellis and Stimson (2009). Table 2 presents correlation coefficients between these measures. The Kim and Fording (2003) measure for the US median voter also builds on the CMP statistics but does not correct for stochastic error and applies a ratio method instead of a subtractive method to tabulate party positions. The ANES measure of party identification is a bi-annual survey that gauges whether respondents think of themselves as Democrat, Independent, or Republican on a 7-point scale. The ANES liberal/conservative measure is a bi-annual survey that gauges whether respondents think of themselves as liberal, moderate, or conservative

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<sup>4</sup>The appendix also provides the empirical analysis using the US median voter data without interpolating non-election values. Doing so reduces the number of observations, but the principal claims of this paper remain equally significant.

on a 7-point scale. The Ellis and Stimson (2009) measure is a compiled and integrated time series of historical and more recent self-identification surveys on whether respondents think of themselves as liberal or conservative.

Table 1: Comparison with alternative voter ideology measures

Name	Availability	N	Range	Mean	Std Dev	Min	Max
<i>US Median Voter</i> (De Neve 2009b)	1920-2008	89	-100/100	8.2	6.6	-7.8	20.9
<i>US Median Voter</i> (Kim and Fording 2003)	1945-2003	59	0/100	48.9	9.5	35.5	66.1
<i>Party identification</i> (ANES)	1952-2008	28	1/7	3.6	0.2	3.3	4.0
<i>Lib/Cons identification</i> (ANES)	1972-2008	18	1/7	4.3	0.1	4.1	4.5
<i>Lib/Cons identification</i> (Ellis and Stimson 2009)	1937-2006	70	0/100	39.3	4.2	31.8	46.3

Table 2: Alternative voter ideology measures correlations table

Name	(1)	(2)	(3)	(4)	(5)
(1) <i>US Median Voter</i> (De Neve 2009b)	1.00				
(2) <i>US Median Voter</i> (Kim and Fording 2003)	0.60	1.00			
(3) <i>Party identification</i> (ANES)	0.77	0.66	1.00		
(4) <i>Lib/Cons identification</i> (ANES)	0.08	0.40	0.49	1.00	
(5) <i>Lib/Cons ID</i> (Ellis and Stimson 2009)	0.51	0.74	0.56	0.61	1.00

*Note:* To facilitate interpretation, the Kim and Fording (2003) and Ellis and Stimson (2009) measures were inversed to obtain data that would also increase when indicating a rise in conservatism.

Of course, in order to measure ideological change in the US one could simply take the variation in electoral success between Democrat and Republican candidates over time. To do so, however, would be a mistake as it would falsely assume that the ideological position of either party has not altered over time. Combining electoral success with an in-depth

analysis of party documents since 1920 allows for a sophisticated measure of voting behavior and ideological change that incorporates voter and party dynamics.

## 2.2 The United States, 1920-2008

For the United States only, the election documents going back as far as 1920 have been coded. Moreover, a special effort was made to code the most recent 2008 election documents.<sup>5</sup> The result is a unique view of the evolution of voting behavior in the United States as shown in Figure 1. The ANES party self-identification measure is added on from when it became available (1952), as well as a polynomial trend line to show the general trend in voting behavior between 1920 and 2008 in the US.

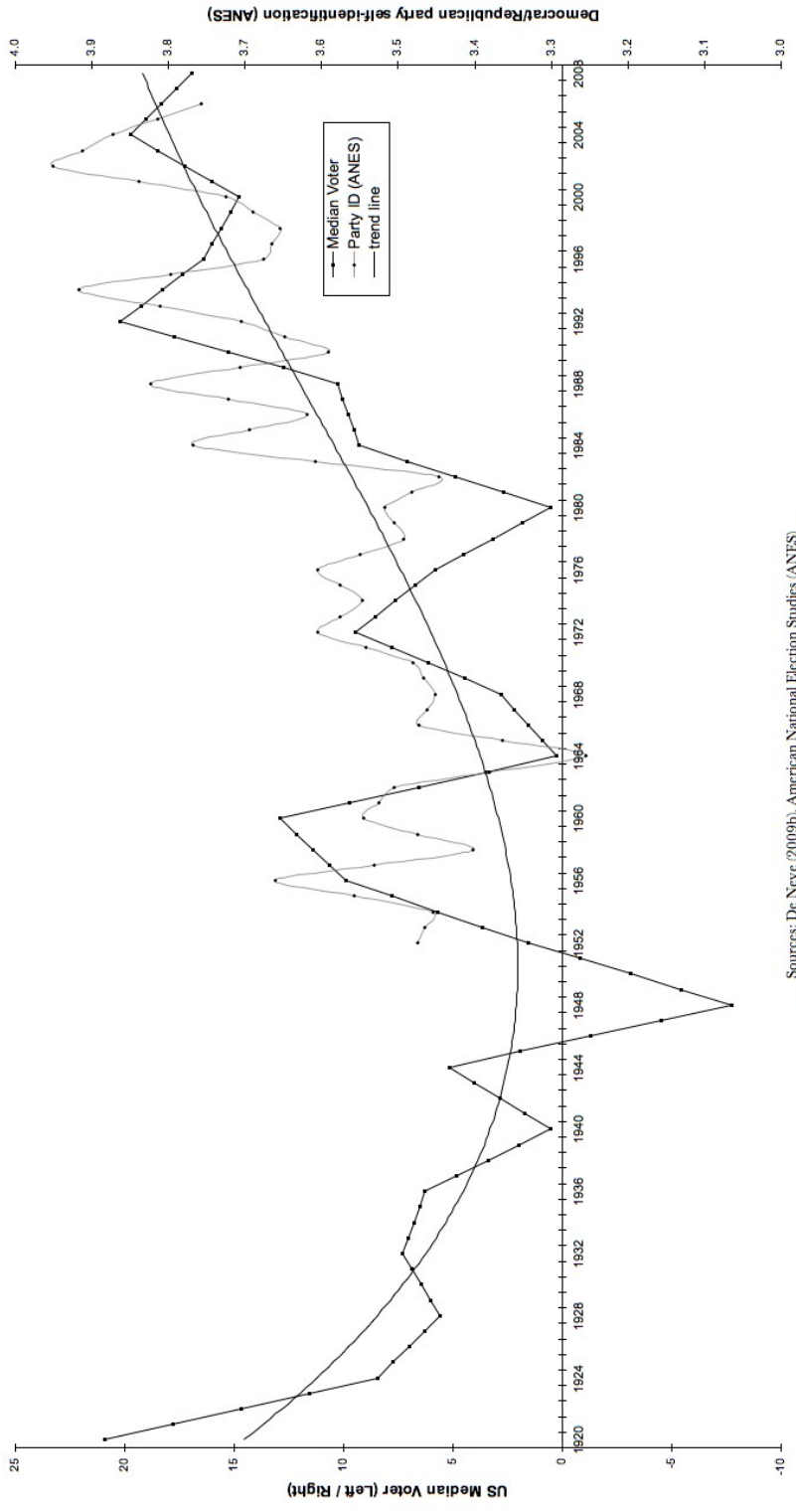
In line with conventional wisdom it shows the US to be generally rightist or conservative. The one time that the US public enters leftist territory is between 1945-50 when, in the wake of the Roosevelt years, Truman finds fertile ground to introduce the Fair Deal that implements a large number of social and economic reforms; including the Housing Act of 1949, an expansion of social security, as well as the first call for universal healthcare. On the international front this less conservative period shows in the large economic aid programs as symbolized by the Marshall Plan. Soon thereafter, however, the US gradually returns to being increasingly more conservative with support for Eisenhower, Nixon, Reagan, and Bush Sr. The elections of Kennedy, Johnson, and Carter do not indicate a turning of the ideological tide though the Kennedy to Johnson and Carter years show a softening conservatism. The mid-eighties see another quantum leap in the conservatism of the American public and culminates with the electoral victory of Bill Clinton who rode the conservative wave on a platform that heralded “the era of big government is over” and promoted fiscal conservatism. The conservativeness of the US electorate drops slightly throughout the second term of the Clinton years and the 2000 election of Bush Jr. However, towards the 2004 re-election of Bush Jr., we note an upswing in conservatism that gradually peels off when we head for the

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<sup>5</sup>Ian Budge (University of Essex) and Judith Bara (QMUL) are to be thanked for their efforts and approval of early release.



Figure 1: The US Median Voter, 1920-2008.



Sources: De Neve (2009b), American National Election Studies (ANES)  
 Note: Median voter data range from -100 (extreme left) to +100 (extreme right).  
 ANES party id range from "strong Democrat" (1) to "strong Republican" (7). Order 3 polynomial trend.

Obama presidency.

What drives these changes in voter sentiment? In his influential contribution, Durr (1993) argued that these broad shifts in ideological sentiment represent responses to changing economic conditions. Before testing this logic empirically, we describe a utility-based model that establishes a relationship between income growth and voting behavior by way of the demand for public goods and the optimal tax rate.

## **3 The Economics of Voting Behavior**

### **3.1 An income growth model**

The logic of the model presented here can be traced back as far as Wicksell (1896) who also considered individuals that allocate their resources between private goods and public goods. The allocation or substitution mechanism for these competing supplies of goods is the rate of taxation and the democratic process allows for a decision. As such, the resulting rate of taxation is chosen by the decisive or median voter who maximizes utility derived from private and public goods given an income constraint. This voter utility model is similar to the textbook model in consumer choice with an individual deriving utility from a convex combination of two bundles of goods subject to a budget constraint that is their personal income. In this voter utility model there are the private goods procured by disposable (post-tax) private income, and there are public goods procured by tax revenue. Public goods are understood in the largest possible sense as comprising all public goods and services provided by government including education, social security, healthcare, utilities, parks, policing, and national defense. This view of public goods is purposely larger than the definition employed by, for example, Meltzer and Richard (1981) who limit government to a redistributive function only.

Given varying marginal utility, individuals will alter their preferred mix of private and public goods to maximize utility with changes in income. This logic was first applied to policy sentiment by Durr (1993) and formalized by Kayser (2009). Both authors assumed

diminishing marginal utility on income and linear marginal utility on public goods. This paper extends variation in policy sentiment to behavior at the voting booth and relaxes the assumption of linear marginal utility on public goods in proposition 2. As the preferred mix of private and public goods changes so does the optimal tax rate that allows for substitution between private and public goods. The way to express preferences on the supply of public goods and the tax rate is to vote for a candidate or party program that would advance those preferences. If the median voter perceives leftist parties as being associated with higher taxation and expansionary government, then a prospering economy with rising disposable income growth will proportionally increase the demand for public goods which translates into more electoral support for the liberal left. Conversely, if the median voter associates parties of the right with lower taxation and smaller government, then a dismal economy that depresses income growth rates will make voters less willing to sponsor public goods and generate more electoral support for the conservative right. It is commonplace to associate leftist parties with higher taxation and expansionary government and parties of the right with lower taxation and smaller government (Downs 1957, Hibbs 1987, Huber and Inglehart 1995, Iversen and Soskice 2006).<sup>6</sup>

In order to formalize a voter preferences model that considers income growth, not levels of income, this paper taps into the literature on reference-dependent utility models that derive utility gains and losses from standard consumption utility models with reference points determined endogenously (Koszegi and Rabin 2006, Koszegi and Rabin 2009). The intuition behind reference-dependent utility models originates in the work of Kahneman and Tversky (1979) who show that the outcome of a choice is not only shaped by absolute values but also by comparison with a reference point. The most obvious economic reference point to which voters will have grown accustomed is income growth in the previous year. We can represent this reference-dependent voter utility model as the sum of the utility gains or losses on

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<sup>6</sup>A different way of relating the logic of this income growth model would be that, while very few actually enjoy paying taxes, in good economic times one is less annoyed with having to part income for the collection of taxes and the provision of public goods. Dire economic times, however, make it particularly painful to pay taxes and will make a campaign promise to reduce taxes all the more appealing.

disposable personal income growth and tax revenue growth.<sup>7</sup> More specifically, with linear utility on public goods,

$$U_v = [(1 - \tau) \Delta y]^\alpha + \tau \Delta y (1 - \delta) \mid \alpha, \delta, \tau < 1$$

where the first term captures the utility gains from income growth ( $\Delta y$ ) after taxes ( $\tau$ ) or disposable personal income growth with which to procure private goods. The second term represents the utility gains derived from public goods that are financed by tax revenue. A parameter to capture the inefficiency of taxation is included ( $1 - \delta$ ). Income growth is subject to diminishing marginal utility, hence the first term is raised to a fractional exponent ( $\alpha$ ).<sup>8</sup>

Solving the first-order condition for the optimal level of taxation ( $\tau^*$ ) gives the below proposition with proof in appendix.

**Proposition 1.** *With linear marginal utility of public goods, the optimal level of taxation increases with rising income growth as given by*

$$\tau^* = 1 - \frac{\alpha^{\alpha-1} (1 - \delta)^{\frac{1}{\alpha-1}}}{\Delta y}$$

When income growth  $\Delta y$  rises, the optimal level of taxation  $\tau^*$  also rises. This optimization result on taxation nicely illustrates the above logic and provides a dynamic link for the relationship between changing economic conditions and policy sentiment that Durr (1993) first described.

When we release the assumption of linear marginal utility on public goods and raise the second term to a fractional exponent ( $\varepsilon$ ) we obtain the following voter utility model<sup>9</sup>

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<sup>7</sup>Of course, the absolute level of personal income will play its role in generating utility to voters. The research presented in this paper, however, is pre-occupied with the explanatory power of income growth rates rather than absolute income levels. Comparable work in political economy typically looks at growth rates when considering personal income, GDP, unemployment, or inflation, instead of absolute levels of those macroeconomic indicators.

<sup>8</sup>This voter utility model is different from the modeling effort by Kayser (2009) in that it dispenses with the arbitrary weighting parameter on public goods and, perhaps more importantly, it is expressed in terms of income growth by relying on the logic of reference-dependent utility models.

<sup>9</sup>This voter utility model is different from the modeling effort by Kayser (2009) in that it allows for an

$$U_v = [(1 - \tau) \Delta y]^\alpha + [\tau \Delta y (1 - \delta)]^\varepsilon \quad | \quad \delta, \tau, \alpha, \varepsilon < 1$$

Solving the first-order condition for the optimal level of taxation ( $\tau^*$ ) and the implicit derivative ( $\frac{d\tau}{dy}$ ) gives the following proposition with proof in appendix.

**Proposition 2.** *If the marginal utility of public goods is sufficiently greater than the marginal utility of disposable income growth ( $\varepsilon > \alpha$ ) then the optimal level of taxation increases with rising income growth as given by*

$$\frac{\frac{\alpha^2}{\varepsilon^2(1-\delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon}-1}}{\{(1-\alpha)\tau - (1-\varepsilon)(1-\tau)\} \frac{\tau^{\varepsilon-2}(1-\tau)^{\alpha-2}}{[(1-\tau)^{\alpha-1}]^2}}$$

The sign on  $\frac{d\tau}{d\Delta y}$  depends on the term  $\{(1-\alpha)\tau - (1-\varepsilon)(1-\tau)\}$ , hence

$$\begin{aligned} \frac{d\tau}{d\Delta y} &> 0 \Leftrightarrow (1-\alpha)\tau - (1-\varepsilon)(1-\tau) > 0 \\ \alpha < 1, \varepsilon < 1 &\Leftrightarrow \frac{(1-\alpha)}{(1-\varepsilon)} \frac{\tau}{(1-\tau)} > 1 \end{aligned}$$

For this condition to hold it is sufficient that  $\varepsilon > \alpha$  as long as  $\tau > 1/2$ . As  $\tau$  decreases, the ratio  $\varepsilon/\alpha$  needs to increase for  $\frac{d\tau}{d\Delta y}$  to remain positive. The case where marginal utility of public goods ( $\varepsilon$ ) is greater than the marginal utility of disposable income growth ( $\alpha$ ) would be common. First, because personal income precedes the possibility of financing public goods, those public goods will be at an earlier stage on their utility function relative to income. Second, demand for public goods and services has been shown to be income elastic, a regularity known as Wagner's Law (Lamartina and Zaghini 2008).<sup>10</sup>

exponential utility function on both income growth and public goods. This model also dispenses with the arbitrary weighting parameter on public goods and is expressed in terms of income growth.

<sup>10</sup>Alternatively, as Stevenson (2001) observes, a slightly different logic from the above would lead to the same conclusion. If voters perceive leftist policies as luxury goods relative to rightist policies, then an individual will maximize her utility by advancing leftist policies when her personal income rises and

## 3.2 Income growth, taxation, and voting behavior

Figure 2 plots the key variables of this voter utility model over time. Variable definitions and sources are available in the appendix. Disposable income growth and the tax rate track each other well, as hypothesized in the above model. A continued rise in income growth rates through the early 1980s is accompanied by a similar rise in the median tax rate. This trend reverses sharply after 1985 with income growth values and the tax rate either dropping or leveling. As income growth contracts, the median voter reevaluates the marginal utility derived from personal income and becomes less supportive of taxation and public spending. Following the above logic, this dynamic leads us to expect that the ideological center of the US electorate will have shifted rightward from the mid-eighties onwards as income growth depresses.

Adding the median voter data visualizes the principal claim of this paper; that ideological change is endogenous to variation in income growth rates. As Figure 2 and the polynomial trend lines show, the ideological center of the electorate joins income growth rates in a long-run trend. While income growth rates increase the electorate becomes more liberal, or to put it more appropriately in the case of the US, the electorate votes less conservative. Conversely, when income growth rates stagnate or contract, voting behavior turns more conservative.

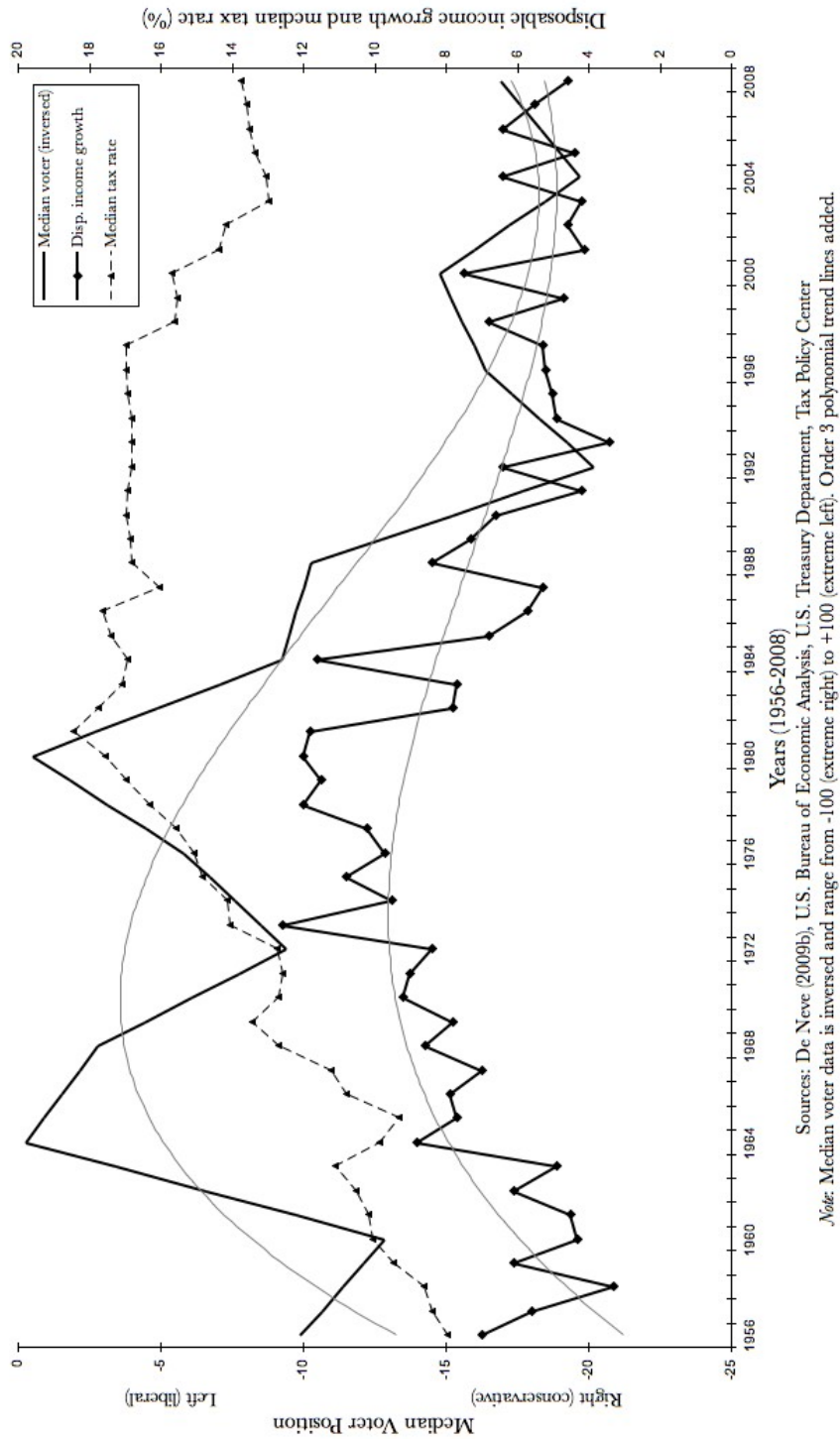
## 3.3 The income growth model versus the redistributive model

It is worthwhile observing that the above voter model is a departure from influential work by, for example, Romer (1975), Meltzer and Richard (1981), Alesina and Rodrik (1994), and McCarty, Poole and Rosenthal (2006). In essence, they model taxation and public spending as the result of redistributive demands by below median income voters. As Pontusson and Rueda (forthcoming) also observe, virtually all of the voter models in political economy have

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rightist policies when income decreases. This logic circumvents the assumption that leftist parties and their policies have to be perceived as generating more expensive public goods as compared to rightist policies. An assumption that could be empirically problematic given that spending priorities by conservative governments have, at times, led to greater levels of public spending than liberal governments (Galbraith 2008).

Figure 2: The Economics of Voting Behavior: Income Growth, Taxation, and Voting Behavior



taken the redistributive model as a point of departure. A dynamic read of these models implies, in the words of McCarty et al. (2006), that “[I]n a responsive democracy, more inequality should typically lead to more redistribution.” In terms of ideological change this would mean that increasing inequality moves the median voter further left in order to drive support for higher levels of taxation and redistributive spending. The logic of the income growth model presented in this paper implies the exact opposite conclusion: a leftist policy agenda will be most successful when rising income growth is equally spread. A rising economic tide that lifts all boats is projected to have greater impact on the demand for public goods than income growth that is limited to few. It is left to future research to extend this analysis cross-nationally and evaluate whether different levels of inequality stimulate or soften the effect of income growth on voting behavior. A preliminary conjecture would be that, *ceteris paribus*, societies with less inequality will have seen a steeper rise in leftist voting as income growth accelerated through the eighties and a sharper conservative turn as income growth decelerated from the mid-eighties onwards. The opposite logic is suggested for societies with high levels of inequality where income growth is less widespread and will thus generate a softened effect on voting behavior.

Besides the opposed implications, two other important elements set this income growth model apart from the prevalent redistributive model. First, as mentioned before, the function of the state is defined as providing public goods and services with the tax revenue it collects. These are understood in the broadest possible sense. The redistributive voter model limits government to a redistributive function and voter’s minds are pre-occupied with calculating whether they will be a net contributor or recipient at the end of the redistributive process. As such, in the most basic version of the Meltzer-Richard model, below median-income voters would have an incentive to support 100% taxation whereas the above median-income voter would prefer no taxation. Second, the redistributive model of economic voting behavior suggests a two-step logic where an individual first has to part with money at tax time before being returned funds through redistributive processes at a later stage. The logic of the



income growth model presented here does not require a two-step investment logic with a possible return of funds. Instead, this model suggests a straightforward spending approach that requires but that voters perceive that increased taxation is positively correlated with improving levels of public goods and services.

Notwithstanding a dramatic rise in inequality in the US over the past three decades, the US voting public has become more conservative by any ideological measure. The redistributive model failed to predict this conservative trend in voting behavior. Many scholars set out to explain this empirical conundrum, coined the “paradox of redistribution” or “Robin Hood paradox,” by pointing to political pressures from the top outweighing pressure for redistribution from the bottom as low-income citizens are less likely to vote or are not entitled to vote (McCarty et al. 2006, Barnes 2007, Pontusson and Rueda forthcoming). A number of scholars offered alternative models to explain pressures for redistribution and partisanship by focussing in on electoral systems, coalitions, and organized labor (Bradley, Huber, Moller, Nielsen and Stephens 2003, Iversen and Soskice 2006, Iversen and Soskice 2009). Others reasoned that demand for insurance rises with income (Moene and Wallerstein 2003), and still others put forth that “our goal is to rescue the idea that income inequality is not only shaped by politics, but also shapes politics” (Pontusson and Rueda forthcoming).

The below empirical analysis gauges the influence of income growth rates on voting behavior and includes an updated measure of inequality by Piketty and Saez (2003). Contrary to the intuitive logic of the Meltzer-Richard redistributive model, the results indicate that public spending does not “result from the difference between the distribution of votes and the distribution of income” (Meltzer and Richard 1981). Variation in income growth rates, however, does provide explanatory power for the demand for public spending and partisanship in the US.

## 4 Empirical analysis

### 4.1 Reverse causality

Very little research on economic voting has attempted to deal with the possibility that voting may also impact economics. Wlezien, Franklin and Twiggs (1997) first raised the endogeneity issue and showed that the estimated effects of economic perceptions on vote choice were substantially less than originally produced in the widely cited work by Lewis-Beck (1988). If voting behavior influences the economic variables then using OLS estimations will inflate the economic effect on voter ideology. Benign neglect of this reverse causality may have increased our understanding of how economic realities affect voting behavior but statistical analyses may be biased. Not unlike the proverbial chicken and egg problem, is there a serious risk to attribute ideological change to variation in economic performance if, in fact, ideological change and the resulting voting behavior may have influenced economic performance in the first place. If so, this reverse causality poses an important endogeneity problem that econometric studies in economic voting appear to have neglected.

The use of one or more valid instrumental variables in a two-stage least squares (2SLS) regression allows for the estimation of causal relationships in the presence of endogenous explanatory variables. The instrument cannot be correlated with the dependent variable (exclusion) but should be highly correlated with the endogenous explanatory variable for which it instruments (relevance). If so, then a 2SLS regression allows for consistent estimation (Acemoglu, Johnson and Robinson 2001, Heckman 2008). This paper proposes to use the measure for net income growth per head in the OECD (minus the US) as instrument for the variation in US disposable income growth. With increasing economic interdependence, a great number of scholars have established that domestic economic performance moves in sync with international economic fluctuations, or that international business cycles may actually induce covariation across domestic business cycles (Campbell and Mankiw 1989, Kose, Otrok and Whiteman 2003). Hence, it is expected that net income growth rates for the OECD

will be strongly correlated with income growth in the US. Kayser (2009) takes this logic one step further and shows that international business cycles influence domestic economic performance which, in turn, induces domestic electoral choice. As such, Kayser concludes that the cross-national success of left and right parties over time is best characterized by “partisan waves” that originate from international economic sources.

The validity of the 2SLS results relies on the assumption that international business cycles have an *indirect* effect on US voting behavior by way of US domestic business cycles. A screening of the first-stage results of the 2SLS models shows that economic indicators for the OECD (minus the US) are relevant instruments for their US domestic counterparts. In fact, net income growth per head in the OECD (minus the US) correlates at 0.88 with US disposable income growth figures, and the first-stage regression coefficients on OECD income growth, and their first differences, obtain p-values below 0.01. The instrumented analysis produces an F-statistic of 5.33 (Stock and Yogo 2005).

## 4.2 Empirical model

Having seen in Figure 2 that ideological change and disposable income growth appear to move in sync over time, we can now turn to a statistical analysis that considers reverse causality and controls for other possible determinants of voting behavior.

Augmented Dickey-Fuller unit-root tests suggest the presence of non-stationary or integrated variables. In similar fashion to Durr (1993) and Kayser (2009), we turn to a general error correction model (ECM) as the most appropriate way to proceed. Practically, the ECM is given by:<sup>11</sup>

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<sup>11</sup>Theoretically, the ECM is given by:

$$\Delta Y_t = \alpha + \beta \Delta X_t + \delta(Y_{t-1} - X_{t-1}\gamma) + \epsilon_t$$

where the error correction term measures the distance from equilibrium that variables move following changes and the time it takes to return to equilibrium. For more info on the ECM please refer to DeBoef and Keele (2008).

$$\Delta Y_t = \alpha + \beta_1 Y_{t-1} + \beta_2 \Delta X_t + \beta_3 X_{t-1} + \beta_4 \Delta Z_t + \epsilon_t$$

where  $\Delta Y_t$  is the dependent variable, ideological change in the position of the US median voter, in year  $t$ .  $\Delta Y_t$  is regressed on its lagged level  $Y_{t-1}$ .  $\Delta Y_t$  is also regressed on both the first difference and the lagged level of the co-integrated independent variable  $X_t$ , in this case US disposable personal income growth rates. A one year lag for income growth rates is used and a 4-year lag of the median voter is used to capture dynamics that span another election.  $Z_t$  are additional controls, here a measure for inequality<sup>12</sup> and US defense spending as a percentage of GDP<sup>13</sup>. The appendix reproduces the analysis using additional controls for inflation, voter turnout, and presidential approval rates. Variable definitions and sources are also listed in the appendix.

### 4.3 Results

Table 3 shows the results of ordinary and instrumented error correction models that test the hypothesis that variation in personal income growth predicts ideological change. Both the OLS and 2SLS models show that increases in disposable personal income growth rates are significantly associated with more leftist or liberal voting behavior in the US. This is the case for the one-year lagged income growth rates (OLS  $p = 0.013$  and 2SLS  $p = 0.024$ ), as well as for the first differences that capture short-run effects (OLS  $p = 0.017$  and 2SLS  $p = 0.020$ ). These results provide strong support for the logic behind the income growth model: an increase in personal income growth rates generates a relative increase in the demand for

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<sup>12</sup>A measure for inequality by Piketty and Saez (2003) is included to control for its possible influence on voting behavior as suggested by, among others, Meltzer and Richard (1981) and more recently by McCarty et al. (2006). The latter authors also show that inequality and the polarization of the US electorate are tightly correlated ( $r = 0.93$ ), hence including inequality as a control also serves the purpose of covering the political polarization of the US public.

<sup>13</sup>US defense spending as a percentage of GDP is included to capture non-economic factors that may influence ideological change. Durr (1993) used dummies for the Vietnam War, Watergate, Iran hostage crisis, and Iran-Contra affair. Kayser (2009) used troop levels in Europe as a control in his cross-national study.

Table 3: Ordinary and 2SLS error correction models on ideological change in the US (1950-2008).

	$\Delta Median\ voter\ OLS$			$\Delta Median\ voter\ 2SLS$		
	Coefficient	SE	P-value	Coefficient	SE	P-value
$\Delta IncomeGrowth$	-0.22	0.09	0.017	-0.42	0.18	0.020
$IncomeGrowth_{t-1}$	-0.28	0.11	0.013	-0.36	0.16	0.024
$Medianvoter_{t-4}$	-0.32	0.04	0.000	-0.34	0.10	0.001
$Defensespending$	-0.31	0.12	0.012	-0.51	0.44	0.245
$Inequality$	0.12	0.06	0.047	0.09	0.08	0.262
Intercept	2.39	3.43	0.489	5.38	5.53	0.330
N	58			36		
R-squared	0.57			0.33		
F statistic				5.33		

*Note:* Variable definitions are in the appendix. The 2SLS instruments are the first differences and a one-year lag of disposable income growth rates in the OECD countries minus the US (1970-2008). Regression coefficients with standard errors (SE) and P-values are presented.

public goods and liberal policies, resulting in a leftward shift of the US median voter.

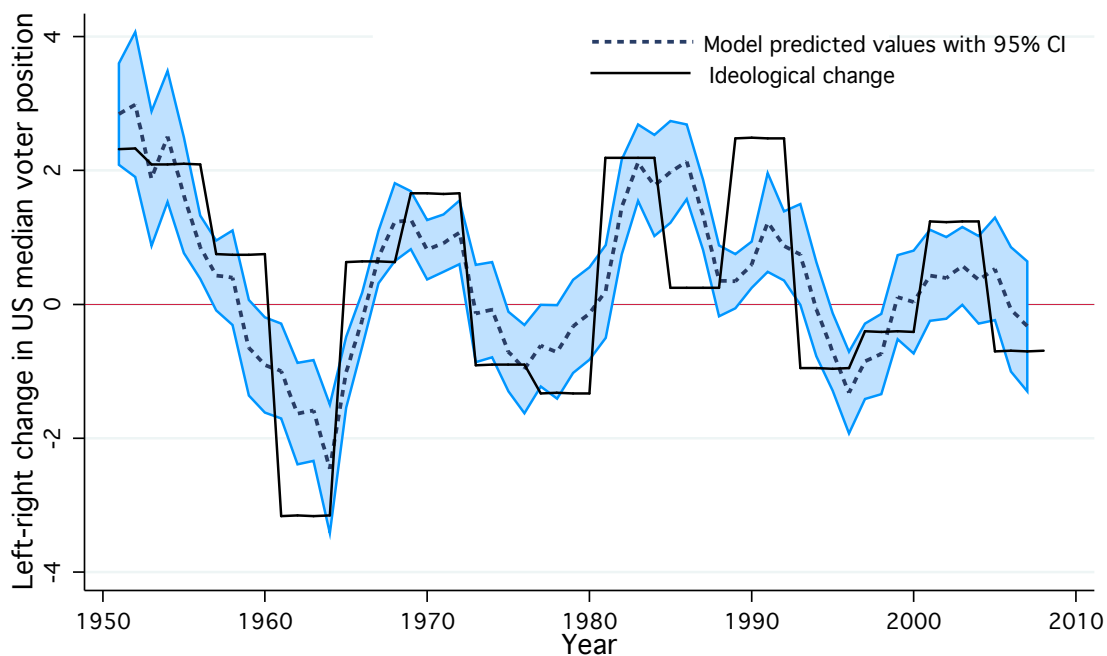
Also of interest is that the measure for inequality and defense spending as a percentage of US GDP come in significant, though both variables lose their significance once we control for reverse causality in the 2SLS model. Note that inequality is positively associated with conservative changes in the position of the US median voter, contrary to the intuitive logic of the Meltzer-Richard model. The 4-year lagged median voter position is negatively associated with current ideological change which would appear to indicate a tendency to somewhat revert ideological course at every new election.

Additional statistical analyses reproduced in the appendix include controls for inflation, voter turnout, and presidential approval rates, as well as median voter data that is not interpolated. The results reported above are robust to these adaptations of the error correction model and coefficients on income growth variables remain significant at the 5% level. Using non-interpolated median voter data reduces the number of observations to fourteen and raises the R-squared to 0.94.

## 4.4 Fitting predicted model values: a retrospective

It is of interest to evaluate the predictive power of our statistical model on past variation in the ideological position of the US median voter. Figure 3 plots the predicted values from the OLS error correction model against actual data for US ideological change. The resulting picture shows that the recursively predicted values track actual ideological change remarkably well; particularly so because the model only integrates a 4-year lag in the median voter position.

Figure 3: Fitting predicted values on US ideological change, 1950-2008



On the basis of 4-year old electoral data and up-to-date information on disposable income, defense spending, and inequality, this model fairly accurately predicts ideological change in the US.<sup>14</sup>

<sup>14</sup>A model built on historical data provides no guarantee for future accuracy. Still, it is an entertaining exercise to contemplate the direction our model would suggest for ideological change in the near future.

## 4.5 Macroeconomic indicators

A unique feature of this paper is that it considers personal income growth statistics instead of the usual macroeconomic suspects that are GDP growth, unemployment, and inflation rates. Income is an obvious fit given that the income growth model presented above hinges on the utility of personal income and the willingness to fund public goods. Still, the few other papers that have elaborated on this logic considered either the usual economic indicators (Stevenson 2001, Markussen 2008, Kayser 2009) or economic expectations (Durr 1993). Table 4 presents the correlation coefficients between the US median voter, income growth, GDP growth, unemployment, and inflation. Income, GDP, and inflation are negatively correlated with the median voter, i.e. a rise in those indicators tends to correlate with a leftist change by the median voter. On the other hand, unemployment is positively correlated with the position of the median voter, indicating that rising unemployment correlates with a more conservative repositioning of the US median voter.

Table 4: Correlations table.

	(1)	(2)	(3)	(4)	(5)
(1) <i>MedianVoter</i>	1.00				
(2) <i>IncomeGrowth</i>	-0.60	1.00			
(3) <i>GDPgrowth</i>	-0.29	0.31	1.00		
(4) <i>Unemployment</i>	0.08	0.15	-0.29	1.00	
(5) <i>Inflation</i>	-0.25	0.56	-0.42	0.41	1.00

*Note:* Variable definitions are in the appendix. Considers the time series between 1950-2008.

Of course, correlation coefficients do not suffice and Table 5 presents the results of ordinary and 2SLS error correction models analogous to Table 3 with GDP growth, unemployment, and inflation as explanatory variables. In line with personal income growth rates, the

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With a deep recession, high unemployment, and slow economic recovery at hand this model would predict the American electorate to have become more conservative, skeptical of public spending, and less inclined to advance public goods such as pro-environment regulation. Hence, unless the Democratic Party or candidate adopts a more rightward policy agenda, we would expect the Republican candidate to pick up more votes in 2012 as compared to 2008.

GDP growth variables in Table 5 are also negatively associated with  $\Delta MedianVoter$  and the lagged GDP growth rate ( $GDPgrowth_{t-1}$ ) is statistically significant. This indicates, once again, that good economic times go hand-in-hand with leftward ideological shifts. This result on GDP growth is verified when introducing the respective OECD data instrumental variables though the low R-squared and F-statistic cast doubt over the validity of this result.

Table 5: Ordinary and 2SLS error correction models on ideological change in the US (1950-2008).

	$\Delta Median voter OLS$				$\Delta Median voter 2SLS$			
	Coeff.	P-value	Coeff.	P-value	Coeff.	P-value	Coeff.	P-value
$\Delta GDPgrowth$	-0.08	0.241			0.08	0.733		
$GDPgrowth_{t-1}$	-0.20	0.027			-0.84	0.077		
$\Delta Unemployment$			0.19	0.277			1.05	0.021
$Unemployment_{t-1}$			-0.21	0.138			0.06	0.849
$\Delta Inflation$			-0.16	0.017			-0.33	0.023
$Inflation_{t-1}$			0.03	0.012			0.02	0.285
$Medianvoter_{t-4}$	-0.27	0.000	-0.29	0.000	-0.33	0.019	-0.26	0.006
$Defensespending$	-0.06	0.485	-0.17	0.132	-0.41	0.509	-0.20	0.650
$Inequality$	0.22	0.000	-0.16	0.259	0.24	0.009	-0.10	0.621
Intercept	-4.53	0.024	8.96	0.085	-1.26	0.803	6.49	0.384
N	58		58		36		37	
R-squared	0.56		0.61		0.00		0.48	
F statistic					0.97		1.95	

*Note:* Variable definitions are in the appendix. The instrumental variables used are the first differences and 1-year lags for GDP growth, unemployment, and inflation in the OECD countries minus the US (1970-2008). Regression coefficients with P-values are presented.

Another material result that survives the instrumented analysis is that the first difference in unemployment ( $\Delta Unemployment$ ) positively predicts US ideological change towards being more conservative. This finding echoes the key result in Kayser (2009) that increases in unemployment are associated with increases in vote intention for conservative parties, especially in “high-clarity” countries such as the US. This unemployment effect does not hinge on the partisanship of the incumbent government and interacts well with the logic of



the income growth model. While it may be surprising to find that rising unemployment actually reduces the demand for expansionary government; given that rising  $\Delta Unemployment$  is negatively correlated with personal income growth ( $r = -0.22$ ), the income growth model rightly predicts a new appreciation of disposable personal income and less willingness to sponsor costly government programs.

Mixed signals are present when looking at inflation. For alignment with the voter utility model we would predict that because rising inflation lowers the purchasing power of personal income it would therefore push voting behavior towards more conservative grounds. Table 5 returns a negative coefficient for  $\Delta Inflation$  and a positive coefficient for  $Inflation_{t-1}$ .

## 5 Conclusion

The US median voter tends to advance a more liberal policy agenda when economic times are good and turns more conservative when the economic situation deteriorates. Error correction models on disposable income growth show it to be a strong predictor of ideological change even after controlling for other possible sources of voting behavior (*Defensespending*, *Inequality*, *VoterTurnout*, *PresidentApproval*) and reverse causality.

That income growth influences voting behavior can be easily understood if one considers a reference-dependent utility model that is similar in set-up as the standard consumer choice model with competing goods. Given diminishing marginal utility on income growth, individuals will alter their preferred mix of private and public goods to maximize utility with changes in income. If the median voter perceives leftist parties as being associated with expansionary government, then a prospering economy with rising disposable income growth will proportionally increase the demand for public goods which translates into more electoral support for the liberal left. Conversely, if the median voter associates parties of the right with smaller government, then a dismal economy that depresses income growth will make voters less willing to sponsor public goods and generate more electoral support for the conservative right. The optimization results of the utility-based model presented in this

paper provide a dynamic link for the fundamental relationship between changing economic conditions and policy sentiment that Durr (1993) first described. The empirical tests leave no doubt about the strength of that relationship.

The alignment of theory with empirical realities is what has plagued the Meltzer-Richard redistributive model (1981) from the start. Still, the intuitive appeal of voting behavior as a function of redistributive demands and inequality has turned their model into the predominant starting point when considering the economics of voting behavior. The strong empirical results of the income growth model and logic presented in this paper may help develop an alternative understanding of voting behavior.

The general lesson to be drawn from this work is that ideological change in the US is endogenous to variation in disposable income growth rates and other measures of economic performance. This result nicely complements the finding in De Neve (2009a) that the left-right ideological position of a country is endogenous to the variety of economic institutions that make up their respective political economy. With both the ideological position and ideological change endogenous to economic realities, it is impossible to deny the pervasive and long-run impact of economics on voting behavior.

## Appendix

### Proof of proposition 1

*With linear marginal utility of public goods, the optimal level of taxation increases with rising income growth.*

$$U_v = [(1 - \tau) \Delta y]^\alpha + \tau \Delta y (1 - \delta) \quad | \quad \alpha, \delta, \tau < 1$$

FOC

$$\frac{dU_v}{d\tau} = (1 - \delta)\Delta y - \alpha \Delta y((1 - \tau)\Delta y)^{\alpha-1} = 0$$

$\Rightarrow$

$$\alpha \Delta y((1 - \tau)\Delta y)^{\alpha-1} = (1 - \delta)\Delta y$$

If  $\Delta y \neq 0, \alpha \neq 0,$

$$((1 - \tau)\Delta y)^{\alpha-1} = \frac{(1 - \delta)}{\alpha}$$

$\Rightarrow$

$$(1 - \tau)^{\alpha-1} = \frac{(1 - \delta)}{\alpha \Delta y^{\alpha-1}}$$

$\Rightarrow$

$$\tau = 1 - \left[ \frac{(1 - \delta)}{\alpha \Delta y^{\alpha-1}} \right]^{\frac{1}{\alpha-1}}$$

$\Rightarrow$

$$\tau^* = 1 - \frac{\alpha^{\alpha-1}(1 - \delta)^{\frac{1}{\alpha-1}}}{\Delta y}$$

## Proof of proposition 2

If the marginal utility of public goods is sufficiently larger than marginal utility of disposable income growth ( $\varepsilon > \alpha$ ) then the optimal level of taxation increases with rising income growth.

$$U_v = [(1 - \tau) \Delta y]^\alpha + [\tau \Delta y (1 - \delta)]^\varepsilon \quad | \quad \delta, \tau, \alpha, \varepsilon < 1 \text{ and } \alpha < \varepsilon$$

FOC

$$\frac{dU_v}{d\tau} = -\alpha \Delta y [(1 - \tau) \Delta y]^{\alpha-1} + \varepsilon \Delta y (1 - \delta) [\tau \Delta y (1 - \delta)]^{\varepsilon-1} = 0$$

$$\begin{aligned} \frac{\tau^{\varepsilon-1}}{(1 - \tau)^{\alpha-1}} &= \frac{\alpha \Delta y^\alpha}{\varepsilon [(1 - \delta) \Delta y]^\varepsilon} \\ &= \frac{\alpha}{\varepsilon (1 - \delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon}} \end{aligned}$$

Implicit derivative

$$F(\tau, \Delta y) = \frac{\tau^{\varepsilon-1}}{(1 - \tau)^{\alpha-1}} - \frac{\alpha}{\varepsilon (1 - \delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon}} = 0$$

$$\frac{d\tau}{d\Delta y} = -\frac{F_{\Delta y}}{F_\tau} = -\frac{-\frac{\alpha^2}{\varepsilon^2 (1 - \delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon} - 1}}{\frac{(\varepsilon - 1) \tau^{\varepsilon-2} (1 - \tau)^{\alpha-1} - (\alpha - 1) (1 - \tau)^{\alpha-2} \tau^{\varepsilon-1}}{[(1 - \tau)^{\alpha-1}]^2}}$$

$$= -\frac{-\frac{\alpha^2}{\varepsilon^2 (1 - \delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon} - 1}}{\{(\varepsilon - 1) (1 - \tau) - (\alpha - 1) \tau\} \frac{\tau^{\varepsilon-2} (1 - \tau)^{\alpha-2}}{[(1 - \tau)^{\alpha-1}]^2}}$$

$$= \frac{\frac{\alpha^2}{\varepsilon^2 (1 - \delta)^\varepsilon} \Delta y^{\frac{\alpha}{\varepsilon} - 1}}{\{(1 - \alpha) \tau - (1 - \varepsilon) (1 - \tau)\} \frac{\tau^{\varepsilon-2} (1 - \tau)^{\alpha-2}}{[(1 - \tau)^{\alpha-1}]^2}}$$

## Descriptive statistics

Table 6: Variable properties

Variable	N	Mean	Std Dev	Min	Max
<i>MedianVoter</i>	89	8.23	6.54	-7.77	20.91
<i>IncomeGrowth</i>	59	7.09	2.52	2.2	12.6
<i>GDPgrowth</i>	59	3.41	2.31	-1.9	8.7
<i>Unemployment</i>	59	5.61	1.46	2.9	9.7
<i>Inflation</i>	59	3.80	3.05	-2.08	13.91
<i>OECD.Income</i>	37	6.35	2.58	2.35	11.66
<i>OECD.GDPgrowth</i>	38	7.00	2.93	3.02	12.86
<i>OECD.Unemployment</i>	44	5.86	2.05	2.57	8.97
<i>OECD.Inflation</i>	38	6.83	3.40	2.37	14.86
<i>Defensespending</i>	59	6.45	2.76	3.00	14.2
<i>Inequality</i>	91	39.22	5.37	32.31	49.74
<i>Turnout</i>	61	55.71	4.11	48.9	62.8
<i>PresidentApproval</i>	59	55.78	13.35	23	84
<i>Taxrate</i>	54	14.06	3.02	7.35	18.44

Table 7: Variable definitions and sources

Variable	Definition and source
<i>MedianVoter</i>	The ideological center of the electorate; ranges from -100 (extreme left/liberal) to +100 (extreme right/conservative). Source: De Neve (2009b).
<i>IncomeGrowth</i>	The annual percent change in US disposable personal income in current dollars (code A067RP1). Source: Bureau of Economic Analysis.
<i>GDPgrowth</i>	The annual percent change in US real gross domestic product. Source: Bureau of Economic Analysis.
<i>Unemployment</i>	The yearly US unemployment rate from the employment status of the civilian noninstitutional population. Source: Bureau of Labor Statistics.
<i>Inflation</i>	The inflation rate or the annual percent change in the Consumer Price Index (CPI). Source: Bureau of Labor Statistics.
<i>OECD.Income</i>	The annual percent change in net national income per head for the OECD minus the US (in current dollars). Source: OECD.
<i>OECD.GDPgrowth</i>	The annual percent change in real gross domestic product (in current dollars) for the OECD minus the US. Source: OECD.
<i>OECD.Unemployment</i>	The yearly unemployment rate in the OECD minus the US. Source: OECD.
<i>OECD.Inflation</i>	The inflation rate or the annual percent change in consumer prices in the OECD minus the US (base year = 2005). Source: OECD.
<i>Defensespending</i>	The annual national defense outlays as a percentage of GDP. Source: OMB (The Budget for Fiscal Year 2009, Historical Tables).
<i>Inequality</i>	The top-10 percent fractile income share (including capital gains) in the US. Source: Piketty and Saez (2003).
<i>Turnout</i>	Voter turnout rate at presidential elections based on voting-age population. Source: McDonald and Popkin (2001).
<i>PresidentApproval</i>	Presidential approval rate at start of each year. Source: Gallup.
<i>Taxrate</i>	The average and marginal combined federal income and employee social security and medicare (FICA) tax rate for four-person families at the same relative positions in the income distribution (median income). Source: Treasury Department (through 1996); Tax Policy Center (1997-2008).

Table 8: Additional OLS and 2SLS error correction models on ideological change in the US (1950-2008).

	$\Delta Median\ voter$				$\Delta Median\ voter\ non-interpolated$	
	OLS		2SLS		OLS	
	Coeff.	P-value	Coeff.	P-value	Coeff.	P-value
$\Delta IncomeGrowth$	-0.22	0.026	-0.42	0.017	-3.00	0.002
$IncomeGrowth_{t-1}$	-0.28	0.054	-0.43	0.068	-4.02	0.006
$Medianvoter_{t-4}$	-0.33	0.000	-0.32	0.005	-1.96	0.000
$Defensespending$	-0.33	0.050	-0.39	0.417	-4.03	0.002
$Inequality$	0.13	0.038	0.09	0.288	-0.41	0.187
$Turnout$	0.01	0.926	0.03	0.802	0.05	0.864
$PresidentApproval$	0.01	0.384	0.01	0.520	0.10	0.149
$Inflation$	0.02	0.830	0.07	0.559	-0.21	0.577
Intercept	1.01	0.839	2.85	0.717	82.43	0.018
N	58		36		14	
R-squared	0.58		0.37		0.97	
F statistic			3.72			

*Note:* The 2SLS instruments are the first differences and a one-year lag of disposable income growth rates in the OECD countries minus the US (1970-2008). There are insufficient observations to run a 2SLS on non-interpolated US median voter data. Regression coefficients with P-values are presented.

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